

DOES FOREIGN AID DISTORT INCENTIVES AND HURT GROWTH?

Theory and evidence from 75 aid-recipient countries

by

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Abstract: In this paper, foreign aid transfers can distort individual incentives, and hence hurt growth, by encouraging rent seeking as opposed to productive activities. We construct a model of a growing small open economy that distinguishes two effects from foreign transfers: (i) a direct positive effect, as higher transfers allow the financing of infrastructure; (ii) an indirect negative effect, as higher transfers induce rent-seeking competition by self-interested individuals. In this framework, the growth impact of aid is examined jointly with the determination of rent-seeking behavior. We test the main predictions of the model for a cross-section of 75 aid-recipient countries between 1975 and 1995. There is evidence that aid has a direct positive effect on growth, which is however significantly mitigated by the adverse indirect effects of associated rent-seeking activities. This is especially the case in recipient countries with relatively large public sectors.

Keywords: Foreign aid. Incentives. Growth.

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1. Introduction

The effects of foreign aid on recipient countries' growth are, at best, ambiguous.¹ In many cases, despite earlier optimistic expectations, aid has failed to boost growth. A prevalent explanation seems to be that foreign aid is often misused and misappropriated. For instance, it can foster corruption in recipient countries.² In this paper, we revisit these issues by studying, both theoretically and econometrically, the joint determination of economic growth and rent-seeking behavior when the driving force is foreign aid.³ The emphasis will be on the distorting effects of aid on private incentives.

We construct a small open economy in which domestic taxes and foreign transfers co-finance public infrastructure. The latter is the engine of long-term growth, as in the model introduced by Barro (1990). We distinguish two effects of foreign transfers on the recipient country's growth rate: (i) A direct positive effect through the financing of public infrastructure. (ii) An indirect negative effect through the distortion of private incentives; the focus will be on rent-seeking competition (modeled as a non-cooperative Nash game) for a fraction of those transfers. We show that under certain conditions (depending on the relation between the magnitude of foreign transfers and the size of the recipient country's public sector), the possibility of extraction from foreign transfers pushes self-interested individuals away from productive work to rent-seeking competition in an attempt to extract resources for their own personal benefit. There is therefore an indirect negative effect, which may counterbalance the manna-from-heaven direct effect that aid typically has on growth. The

¹ There is a rich empirical literature. Empirical studies on the growth impact of aid include Mosley et al. (1987), Boone (1996), World Bank (1998), Svensson (1999), Easterly (2001), Mosley et al. (2004) and Dalgaard et al. (2004). Moreira (2003) provides a useful summary of evidence from cross-country regression studies: out of seventy-two regressions, the estimated effect is positive in forty regressions, non-significant in thirty-one and negative in one regression.

² As Alesina and Weder (2002) point out, foreign aid fosters corruption by increasing the size of resources that interest groups fight over. See also Svensson (2000).

³ The term rent seeking refers to the (both privately and socially) costly pursuit of income and wealth transfers (see Drazen, 2000, p. 335). We will use the terms "rent seeking", "appropriation" and "extraction" interchangeably.

larger the amount of aid and the size of the recipient country's public sector, the higher is the rent seeking effect.

We test these predictions by using data for a pooled cross-section of 75 aid-recipient countries between 1975 and 1995 for which proxies of rent-seeking behavior are available. Following most of the related literature, we use the Knack-Keefer (1995, 1997) indices as measures of rent-seeking behavior. Our econometric results indicate that, when economic growth and rent seeking are examined jointly with aid transfers, a rise in the latter exerts *ceteris paribus* a direct positive effect on economic growth. However, this effect is mitigated by an endogenous increase in rent-seeking activities triggered by the very same rise in aid. Eventually, the net growth effect of aid is substantially smaller compared to the direct positive effect. When we partition recipient countries according to the size of their public sector, and in accordance with our theory, there is evidence that the deleterious effect of aid upon incentives, and in turn growth, is significant only in recipient countries with relatively large public sectors. This finding confirms the intuitive belief that rent seeking and corruption takes place mainly through government activities, as corrupt government officials have some effective property rights over the aid transfers they are allocating (see e.g. Tanzi, 2002).

What is the related theoretical literature? At the heart of the problem of rent seeking, there is a common-pool problem. That is, there is a possibility of common access to a social resource (here, foreign aid), which opens the door to anti-social behavior among self-interested individuals and hurts growth.⁴ A model close to ours is Svensson's (2000). The author develops a game-theoretic model in which rent seekers compete for a fraction of an exogenously given output and shows how a cooperative solution, namely one without rent seeking, can be achieved if the game is repeated over time; foreign aid makes such a

⁴ Different papers differ in what the common-pool resource is. For surveys, see Persson and Tabellini (2000, chapters 7 and 14) and Drazen (2000, chapter 10.7). Some recent related papers, not included in Persson and Tabellini (2000) and Drazen (2000), are: Svensson (2000) where the common-pool is aid transfers; Leite and

cooperative solution more difficult to sustain. Our paper differs mainly because here we use a standard model of growth, in which public infrastructure is co-financed by domestic taxes and foreign transfers, where the latter redistribute from above-average to below-average income countries. We manage to get an analytic and testable solution that distinguishes the direct and indirect effects of aid upon private incentives and economic growth.

What is the related empirical literature? Most of the existing studies have focused on bivariate relations between aid and corruption (see e.g. Alesina and Weder, 2002, Tavares, 2003), growth and aid (see e.g. the references in footnote 1 above) and growth and corruption (see e.g. Mauro, 1995, Knack and Keefer, 1995). Our paper bridges a gap between these empirical relations by estimating the joint determination of rent seeking, growth and aid. To our knowledge, there has been no attempt so far to examine this link. Svensson (1999) has studied the simultaneous determination of growth, aid and a democracy index, while Svensson (2000) has focused on the simultaneous determination of rent seeking and aid without including growth and in particular the direct and indirect effects that aid may have on growth. Recent studies by Dalgaard et al. (2004), who have examined the joint determination of growth, aid and policy, and Mosley et al. (2004), who have examined the joint determination of poverty, aid and policy, have also recognized the possibility that aid may have indirect effects on growth or poverty via the channel of policy. However, they have not considered the distortion of private incentives, and, in particular, how foreign aid affects growth and incentives in a theory-consistent setup.

Our findings underline the need to search for possible channels through which aid affects growth. In particular, our findings may help to explain why past studies were frequently unsuccessful in establishing a substantial positive impact of aid on growth, as aid-induced rent-seeking activities were typically ignored. Our paper may also offer a potential

Weidmann (2002) and Grossman and Mendoza (2003) where the common-pool is natural resources; Mohtadi and Roe (1998, 2003), Mauro (2002) and Park et al. (2003) where the common-pool is government tax revenue.

resolution to the “micro-macro paradox” (Mosley, 1986), according to which aid is found to impact positively when it is evaluated at firms’ investment level by means of cost-benefit project analysis, whereas this positive effect largely evaporates at the level of cross-country regressions. Finally, our paper offers a support to those who argue for conditional aid, where conditionality has to do with the size and the role of the government sector in recipient countries.⁵

The rest of the paper is organized as follows. Section 2 presents a theoretical model. Section 3 specifies the econometric model and describes the data. Section 4 presents the empirical results. Section 5 concludes and discusses policy implications.

2. A theoretical model of growth, incentives and foreign transfers

We will incorporate foreign transfers, as well as rent-seeking competition for a fraction of them, into a model of growth. In particular, we will build on the model introduced by Barro (1990), which is well known and algebraically tractable.⁶

2.1. Informal description of the model

The key features of the theoretical model are as follows: (a) The government uses domestic tax revenues and foreign transfers to finance public productive services. The latter provide a positive production externality to private firms. (b) Only a fraction of foreign transfers is actually used to finance public services. The rest can be extracted by self-interested individuals. Specifically, we assume that households can extract from total foreign transfers to

⁵ Many authors have emphasized that the effectiveness of aid depends on domestic policies. See subsection 4.3 and section 5 below for references and further details.

⁶ The focus is on rent seeking. Thus, many other interesting issues related to aid are left out. For instance, Svensson (2000) studies reputational equilibria as well as the donor’s behavior (but he uses a setup with exogenous output in which rich countries are more prone to rent seeking). Chatterjee et al. (2003) study how aid affects the transitional dynamics of a small open economy (but they abstract from incentives). Our theoretical model is close to that in Mohtadi and Roe (1998, 2003), Mauro (2002) and in particular Park et al. (2003); however, here individuals extract from aid transfers and not from tax revenues. There are many other interesting extensions (some of them are mentioned below) but we deliberately keep the model simple.

increase their own personal wealth.⁷ In doing so, they compete with other households. This rent-seeking competition will be modeled as a non-cooperative (Nash) game among self-interested individuals. (c) Extraction comes at a private cost.⁸ Specifically, it requires time and effort. Thus, each household chooses optimally (in addition to consumption and saving) the allocation of its time and effort to productive work and rent-seeking activities.⁹ (d) The fraction of foreign transfers appropriated by each individual is proportional to the effort and time that he, or she, allocates to rent-seeking competition relative to the total effort and time allocated to rent-seeking competition by all individuals. Then, the aggregate rent-seeking effort will deplete the flow of foreign transfers earmarked for the finance of public services. (e) The economy is small so that it takes the rest of the world as given. For simplicity, we assume that there is only one link with the rest of the world and this is via foreign transfers.

There are $i = 1, 2, \dots, I$ identical households and $i = 1, 2, \dots, I$ identical firms. Concerning quantities, we will distinguish individual variables from their aggregate counterparts by using lower-case and upper-case letters respectively. Thus, for any x ,

$X \equiv \sum_{i=1}^I x^i$ is the aggregate economy-wide variable.

2.2. Firms' behavior

Each firm i maximizes profits, π^i :

$$\pi^i = (1 - \theta)y^i - rk^i - wl^i \tag{1}$$

⁷ We could assume that households also extract from collected tax revenues. This does not change our main results. We could also assume that firms, like households, rent seek. This is not important to our results because households are also firm-owners in this class of models. We could also assume that policy-makers are rent seeking; again this would not be important, as adding more types of self-interested individuals does not affect our main result.

⁸ Breaking the law, bribing, lobbying, extracting favors from the government, etc, are costly activities.

⁹ The idea is as in e.g. Baumol (1990), Murphy et al. (1991) and many others, where individuals decide how to allocate their activities between “productive” ones (such as work, innovation and entrepreneurship) and “unproductive” ones (such as rent seeking and violating property rights).

where $0 < \theta < 1$ is a common proportional output tax rate;¹⁰ y^i , k^i and l^i are i 's output, capital input and labor input; and r and w are the interest rate and wage rate.

At the firm's level, the production function is:

$$y^i = A(k^i)^\alpha (l^i)^{1-\alpha} \left(\frac{G}{I}\right)^{1-\alpha} \quad (2)$$

where $A > 0$ and $0 < \alpha < 1$ are parameters, and G is aggregate public production services.¹¹

Each firm i acts competitively by taking prices, policy variables and aggregate outcomes as given. The first-order conditions for k^i and l^i are simply:

$$r = (1 - \theta)\alpha \frac{y^i}{k^i} \quad (3a)$$

$$w = (1 - \theta)(1 - \alpha) \frac{y^i}{l^i} \quad (3b)$$

so that, with constant returns to scale at the firm's level, profits are zero in equilibrium.

2.3. Households' behavior

Each household i maximizes inter-temporal utility:

$$\int_0^{\infty} \log(c^i) e^{-\rho t} dt \quad (4)$$

where c^i is i 's consumption and $\rho > 0$ is a time discount factor.

At each instant, household i consumes, c^i , saves in the form of an asset, a^i , and allocates one unit of labor time¹² between productive work and rent-seeking activities.

Specifically, let $0 < \eta^i \leq 1$ denote the fraction of i 's time allocated to productive work and

¹⁰ We could use taxes on households' income. The type of distorting taxation used is not important to our results.

¹¹ We assume that it is the average, rather than the total G , that provides production externalities to each firm in order to avoid scale effects in equilibrium. This is not important for the theoretical derivations later on.

¹² For algebraic simplicity, we assume that each household has one unit of labor time available. Exogeneity of total labor time/effort justifies why leisure, or effort time, is not included as an argument in (4).

$0 \leq (1 - \eta^i) < 1$ the fraction allocated to rent seeking. The budget constraint is (a dot over a variable denotes a time derivative):

$$\dot{a}^i + c^i = ra^i + w\eta^i + \frac{(1 - \eta^i)}{I - H} \Delta T \quad (5)$$

where $I - H \equiv \sum_i (1 - \eta^i)$ denotes aggregate rent seeking activities, T is total aid transfers and $0 \leq \Delta < 1$ is the aggregate degree of extraction (see below for its determination). The central idea is that there is a perceived pie ΔT , and each individual tries to extract a fraction of it by competing with all other individuals.¹³

Each household i acts competitively by taking prices, policy variables and aggregate outcomes as given. Combining the first-order conditions for consumption, saving and extraction (c^i, a^i, η^i) , we get:

$$\dot{c}^i = c^i(r - \rho) \quad (6a)$$

$$w = \frac{\Delta T}{I - H} \quad (6b)$$

where (6a) is a standard Euler equation and (6b) implies that net returns from work and appropriative competition are equal in equilibrium.

2.4. The government budget constraint

Assuming a balanced budget, the government budget constraint is:

$$G = \theta Y + (1 - \Delta)T \quad (7)$$

so that public services, G , are financed by domestic taxes, θY , where $Y \equiv \sum_i y^i$, and the

fraction of foreign aid that is not taken away by rent seekers, $(1 - \Delta)T$.

¹³ For a similar way of modeling rent seeking competition, see e.g. Murphy et al. (1991), Svensson (2000), Grossman and Mendoza (2003), Park et al. (2003) and many others. This modeling goes back to Tullock's (1967) original rent-seeking model.

2.5. Decentralized competitive equilibrium

This subsection will solve for a Decentralized Competitive Equilibrium (DCE) in a growing small open economy, for any foreign transfer policy and any feasible domestic policy. Specifically, a DCE is defined to be a Nash equilibrium in individuals' decisions in which: (i) each individual firm and household maximize their own profit and utility respectively by taking the actions of other individuals as given; (ii) all markets clear;¹⁴ (iii) individual decisions are consistent with aggregate decisions;¹⁵ (iv) all constraints, including the economy's resource constraint, are satisfied. For simplicity, we will focus on a symmetric DCE, i.e. in equilibrium private agents (firms and households) are alike ex post. Thus, from now on, the superscript i is omitted.

To get a closed-form solution for the DCE, we will choose - without loss of generality - a convenient specification for foreign aid transfers, T . In particular, we assume:

$$T = \mu(\bar{Y} - Y) \quad (8a)$$

where $\mu > 0$ is a redistribution parameter. According to the redistribution rule in (8a), foreign transfers, T , given to a country are a fraction of the deviation of that country's income, Y , from worldwide average income, \bar{Y} . Thus, if $\bar{Y} > Y$, the country is a recipient; if $\bar{Y} < Y$, the country is a donor (our analysis will be in terms of a recipient country).¹⁶ Since the rest of the world is taken as given, we assume $\bar{Y} = \lambda Y$, where $\lambda > 1$ is a measure of inequality between

¹⁴ That is, $\sum_i l^i = \sum_i \eta^i$ in the labor market and $\sum_i k^i = \sum_i a^i$ in the capital market. As said above, for simplicity, we have assumed away cross-country factor (labor and capital) mobility.

¹⁵ Consistency of individual and aggregate decisions implies that, in equilibrium, the aggregate degree of

extraction, Δ , equals the per capita time spent in rent-seeking activities. Thus, $\Delta = \frac{I-H}{I} \equiv \frac{\sum_i (1-\eta^i)}{I}$, which comprises standard modeling in general equilibrium models with externalities. Note that we use the average time, rather than the total time, in order to avoid scale effects. The adopted specification does not preclude the possibility that Δ depends on other factors in addition to rent seeking activities.

¹⁶ See Park and Philippopoulos (2003) and the references cited therein for similar state-contingent redistributive rules. The rule in (8a) is consistent with several institutional arrangements on foreign transfers.

the domestic economy and the rest of the world.¹⁷ Obviously, λ depends on a number of socio-economic factors, whose specification is an empirical matter (see e.g. Williamson, 1998); we will return to this issue in the empirical section below. In this section, we will solve for a DCE for given λ . Thus, the foreign aid-to-output ratio becomes:

$$\frac{T}{Y} = \mu(\lambda - 1) \equiv \tau \quad (8b)$$

where $\tau \geq 0$ denotes an “effective redistribution” parameter.

We can now solve for a DCE. It is straightforward to show that (1)-(8) give:¹⁸

$$\dot{c} = c \left[(1-\theta)\alpha A^{\frac{1}{\alpha}} \eta^{\frac{1-\alpha}{\alpha}} (\theta + \eta\tau)^{\frac{1-\alpha}{\alpha}} - \rho \right] \quad (9a)$$

$$\dot{k} = [1 + \tau - (\theta + \eta\tau)] A^{\frac{1}{\alpha}} \eta^{\frac{1-\alpha}{\alpha}} (\theta + \eta\tau)^{\frac{1-\alpha}{\alpha}} k - c \quad (9b)$$

$$\eta = \frac{(1-\alpha)(1-\theta)}{\tau} \quad (9c)$$

Equations (9a)-(9c) give the paths of (c, k, η) for any domestic economic policy, as summarized by the tax rate $0 < \theta < 1$, and any foreign transfer, as summarized by the redistribution parameter $\tau \geq 0$.

An advantage of the model is its simplicity. Equation (9c) gives η . If $\tau = (1-\alpha)(1-\theta)$, then $\eta = 1$ so that all effort is allocated to work. If, on the other hand, $\tau > (1-\alpha)(1-\theta)$, then $0 < \eta < 1$ so that only a fraction of effort is allocated to work. In other words, there is rent-seeking activity ($0 < \eta < 1$), only if the foreign aid-to-output ratio (τ)

¹⁷ We set $\lambda > 1$ because the analysis is in terms of a recipient country. Note that the functional specification in (8a)-(8b) does not violate the linear structure of the model, and hence allows us to get a closed-form analytical solution for the competitive equilibrium, as in Barro-type models.

¹⁸ Equation (9a) comes from the Euler equation (6a). Equation (9b) is the economy’s resource constraint, $\dot{k} = y + T - G - c$, and arises from the budget constraints (1), (5) and (7). Equation (9c) follows by equating (3b) and (6b) and using the market-clearing conditions.

and/or the domestic tax rate (θ) are high enough.¹⁹ In turn, having solved for η , (9a) can give

the balanced growth rate, $\gamma \equiv \frac{\dot{c}}{c} = \frac{\dot{k}}{k}$, and (9b) can give the consumption-to-capital ratio, $\frac{c}{k}$.

Inspection of (9a) and (9b) reveals that cases with rent seeking ($0 < \eta < 1$) are associated with a lower growth rate and a lower consumption-to-capital ratio than cases without rent seeking ($\eta = 1$).

We will close with comparative static results by focusing on the effects of the foreign aid-to-output ratio, τ . It is intuitively useful to consider first the benchmark non-interesting case without rent seeking, $\eta = 1$. In this case, η is obviously independent of τ , so that (9a) implies that there is only a direct positive effect from τ on the growth rate, γ , denoted as

$\frac{\partial \gamma}{\partial \tau}^{(+)\text{ direct}}$. Consider now the case with rent seeking, $0 < \eta < 1$. Equation (9c) implies $\frac{\partial \eta}{\partial \tau} < 0$,

i.e. a higher foreign aid-to-output ratio leads to a lower fraction of effort allocated to work

relative to rent seeking. In turn, (9a) implies $\frac{\partial \gamma^{total}}{\partial \tau} = \frac{\partial \gamma}{\partial \tau}^{(+)\text{ direct}} + \frac{\partial \gamma}{\partial \eta} \frac{\partial \eta}{\partial \tau}^{(-)}$. This implies that an

increase in τ exerts two effects on growth, γ : (i) a direct positive effect; (ii) an indirect

negative effect through smaller effort allocated to work, $\frac{\partial \gamma}{\partial \eta} \frac{\partial \eta}{\partial \tau} < 0$. The direct positive effect

arises because foreign aid finances public infrastructure. The indirect negative effect arises

because aid transfers distort the incentives of self-interested individuals. Specifically, the

possibility of extraction pushes them away from productive work to appropriative

competition. Then, as the amount of aid increases, individuals (who do not internalize the

adverse effect of their rent-seeking activities on aggregate outcomes) become more

¹⁹ Conditions for rent seeking like this are usual in the literature. For instance, in models of social conflict, it is optimal to follow appropriative, non-cooperative behavior when assets to be appropriated are relatively high (for a survey, see Drazen, 2000, chapter 10.7). Therefore, following most of the literature, we first assume that it is

aggressive. This is at the society's expense for two reasons: first, it reduces the resources available for public infrastructure (i.e. it increases Δ); second, it reduces effective labor supply (i.e. it reduces η).²⁰

To sum up, when the foreign aid-to-output ratio and/or the domestic tax rate are sufficiently high, we end up in a bad equilibrium with rent-seeking competition. This exerts an indirect negative effect on growth, which works in opposite direction from the direct positive effect that aid typically has on growth.

3. Econometric specification and description of data

This section will develop an empirical framework suitable to test the predictions of the theoretical model (9a)-(9c). The key prediction is that a rise in foreign aid transfers has a direct positive effect on growth by enhancing public infrastructure, but it also triggers an indirect negative effect by pushing individuals in recipient countries to rent-seeking activities. By using a cross-country dataset (see subsection 3.2 below for data description), we will search for these two opposite effects in the context of an econometric model that determines jointly the rate of economic growth, rent seeking activities and foreign aid transfers.

As Svensson (2000) points out, any test of this form is bound to be only suggestive for several reasons. For instance, long-term time series observations are not available for aid and rent seeking, and thus the analysis can only be confined to the medium-term impact of these variables on growth. In addition, rent-seeking activities are hard, if possible at all, to measure. Hence, any empirical methodology can only utilize proxy variables, which can hopefully provide adequate description of this type of activities.

possible for individuals to have access to a common-pool resource (e.g. aid); we then specify the conditions under which this possibility materializes and thus leads to anti-social, rent-seeking activities.

²⁰ The effects of the tax rate (θ) are similar to those of the foreign aid-to-output ratio (τ). See Mohtadi and Roe (1998, 2003), Mauro (2002) and Park et al. (2003) for the effects of government policy in similar models.

3.1. From theory to testing

We will consider linear regressions of the form:²¹

$$\text{growth rate} = G(\text{rent seeking, aid, tax rate; control variables}) \quad (10a)$$

$$\text{rent seeking} = R(\text{aid, tax rate; control variables}) \quad (10b)$$

According to the comparative static properties of (9a)-(9c), and focusing on cases with rent seeking, we expect a positive sign for aid and a negative sign for rent seeking in the growth regression (10a), combined with a positive sign for aid in the rent-seeking regression (10b). Also, a higher domestic tax rate is expected to increase rent seeking activities in (10b), whereas the (Laffer-type curve) effect of the tax rate on growth is ambiguous in (10a).

Equation (8b) suggest that aid transfers (specifically, the foreign aid-to-output ratio) to a particular country increase with its inequality vis-à-vis the rest of the world. Several authors have already pointed out the potential pitfalls associated with the simultaneity bias when aid is treated as an exogenous variable in growth regressions; this accords with the approach adopted by Burnside and Dollar (2000).²² We will therefore assume that aid is jointly determined with growth and rent seeking. In particular, we will use a linear regression of the following form (where a measure of inequality will be defined in subsection 3.2. below):

$$\text{aid} = A(\text{inequality; control variables}) \quad (10c)$$

To control for other possible determinants of the endogenous variables in (10a)-(10c), we follow standard practice and include a number of auxiliary determinants (control variables); their choice is mainly dictated by the studies of Burnside and Dollar (2000),

²¹ Equations (10a) and (10b) follow from (9a) and (9c) respectively. We omit (9b), which gives the consumption-to-capital ratio, because data for capital stocks are not available for most aid-recipient economies. The specification in (10a)-(10b) is similar to that in Leite and Weidmann (2002). The difference is that, in their paper, the driving force of corruption is natural resource abundance.

²² For a survey of the empirical literature, see Hansen and Tarp (2001) who stress that, with a 5-year average data sample (as the one utilized here), treating aid as endogenous, but predetermined, implies that any decisions on the allocation of aid are made on the basis of a 5 to 6 year planning horizon. We maintain this assumption here but we will also report - for comparison purposes - some results when aid is exogenous. As said above, Svensson (1999) has studied the simultaneous determination of growth, aid and a democracy index. Dalgaard et al. (2004) have studied the simultaneous determination of growth, aid and policy. Mosley et al. (2004) have studied the joint determination of poverty, aid and policy.

Svensson (2000), Alesina and Weder (2002) and Persson and Tabellini (2003, chapter 3). These control variables include country characteristics like the initial level of per capita GDP, regional dummies, population size, as well as measures of ethnic conflict and political instability.

Although the inclusion of these control variables is standard, it is useful to provide some brief explanations. First, the initial GDP per capita is used to control for convergence arguments, as well as for recipient countries' needs and motives (see also Svensson, 2000). Second, regional dummies typically capture geographical and historical characteristics of the countries examined; following Burnside and Dollar (2000), dummies for Sub-Saharan countries and East-Asian countries are included in the growth equation. The same regional dummies are used in the rent-seeking equation along with a dummy for Central American countries as in Svensson (2000). In a similar vein, following Burnside and Dollar (2000) the aid equation includes a dummy for Central American countries (which are closely linked to the US), as well as two additional dummies, one for Egypt (an important ally of the US) and one for the Franc zone countries (which are closely linked to France).²³ Third, population size (in logs) is included in the aid equation to capture possible scale effects, as countries with smaller populations are more likely to receive relatively large aid transfers due to donors' strategic interests like voting in international organizations.

Finally, measures of ethnic diversity and political instability are included in all three equations. Ethnic diversity is supposed to capture the ethno-linguistic fractionalization of groups within a country, which is correlated with bad policies and low growth (see Easterly and Levine, 1997), and also takes into account the finding by Svensson (2000) that aid in countries suffering from competing social groups is associated with rent seeking. In turn, political turmoil will be used to capture civil unrest, which is also believed to affect growth

²³ See also Svensson (1999, page 284) on the effect of these variables on growth through aid flows. Persson and Tabellini (2003) use similar regional dummies to capture the effects of colonial history.

(Drazen, 2000). Another variable, usually related to economic conditions and policies, is openness. Measures of all these variables will be defined in subsection 3.2 below.

As far as the estimation method is concerned, most of the literature relies on Instrumental Variable techniques or Panel methods with fixed effects. Taking into account the data availability and the nature of variables utilized, which are largely time-invariant, we will opt for 2SLS estimations in the pooled cross-section (with a core set of instruments consisting of regional dummies, initial GDP, population, government size, as well as measures of ethnic diversity, political instability and openness in extended specifications). This approach aims at capturing the possible endogeneity of rent seeking and aid in the context of growth regressions. In addition, we will report 3SLS results, since the simultaneous estimation of the system (compared to an equation-by-equation estimation method like 2SLS) has the advantage of not imposing any restrictions on the correlation between the error terms, which may improve the estimates on grounds of efficiency in the case of non-zero elements in the variance-covariance matrix. This might be important in the current setup (where, for instance, an unexpected shock in aid can be correlated with the disturbance in the growth equation).

3.2. Description of data

The data used come mainly from three sources: the Penn World Tables, version 6.1 (Heston et al., 2002), the IRIS data set (obtained by *countrydata.com*), and the World Bank database on aid developed by Chang et al. (1998).

The Penn World Tables provide a number of variables. The GDP per capita in constant prices is used to obtain five-year average growth rates (denoted as *growth rate* in the Tables), the log of initial GDP per capita (denoted as *lgdp*) and the log of population (denoted as *lpop*). In the theoretical model, the tax rate is a measure of the size of the government in a particular country. Following most of the literature, we will use the five-year average of

government spending as a share of output to measure the size of government.²⁴ This will give us a five-year average of a variable denoted as *government size*. To get a measure of inequality in the aid regression (10c), we will use the recipient country's log of initial per capita income, *lgdp*, and the log of initial per capita income in high-income OECD countries (following the World Bank classification), denoted as *ldon*. Aid transfers are expected to decrease with the former and increase with the latter, reflecting the idea that as the donors' income increases, aid increases. Finally, we define a country's *openness* as the sum of exports plus imports over GDP to obtain five-year averages.

The source for aid data is the World Bank database on foreign aid. The files included in this database contain the conventional and the adjusted measures of official development assistance to a set of 133 countries between 1975 and 1995. The principal component of this dataset is Effective Development Assistance (EDA), an aggregate measure of aid flows combining total grants and the grant equivalents of all official loans. EDA is computed on a loan-by-loan basis to reflect the financial cost the creditor incurs in making loans on concessional terms. Details on this variable are in Chang et al. (1998). By use of this dataset, we construct five-year averages of the ratio of EDA (in current units relative to current GDP) denoted as *aid*.

To obtain a measure of rent seeking, we use the IRIS dataset (version IRIS-3), which contains annual values for indicators of quality of governance, corruption and violation of property rights over the period 1982-1997, as constructed by Stephen Knack and the IRIS Center, University of Maryland, from monthly ICRG data provided by Political Risk Services.²⁵ This dataset has been used by, among others, Knack and Keefer (1995), Svensson

²⁴ For many developing countries, data on tax revenues are limited. Persson and Tabellini (2003, pp. 49-50) find that tax revenue in a given country in nominal terms is more sensitive than government expenditure to the state of the economy implying a smaller reaction when both measures are scaled to GDP. Also, notice that Tanzi and Schuknecht (2000) use government expenditure as a percentage of GDP, which includes interest payments. However, since this is not available for all non-OECD countries, we use here government spending.

²⁵ Obviously, rent seeking can take a variety of forms other than corruption, bureaucracy or property risks. This type of data, however, is not available at any form. Apart from data non-availability, our choice is also motivated

(2000), Alesina and Weder (2002), Fredriksson and Svensson (2003), Tavares (2003) and Barro and Sala-i-Martin (2004). Following most of the literature, we take the sum of five subjective indices available by the IRIS dataset: namely, “corruption in government”, “rule of law”, “risk of repudiation of government contracts”, “risk of expropriation” and “quality of bureaucracy”. From these indices, “corruption in government”, “rule of law”, and “quality of bureaucracy” range in value from 0 to 6, whereas “risk of repudiation of government contracts” and “risk of expropriation” are scaled from 0 to 10 with higher values indicating better ratings, i.e. less corruption and less rent seeking. We then construct a new variable (measured at a 50-point scale) by converting “corruption in government”, “rule of law”, and “quality of bureaucracy” to a 10-point scale, and then adding them up to the other two indices. The resulting sum is finally averaged for each country for each 5-year period and multiplied by (-1) to give a measure of rent-seeking activities (denoted as *rent seeking*) in the society at large. Finally, from the same dataset, we also use an index of ethnic diversity (denoted as *ethnic*). Following related studies (see e.g. Knack and Keefer, 1995), we also utilize a measure of political instability measured by the average number of political assassinations per million people per year (denoted as *assassinations*).

4. Empirical evidence from pooled cross-section data

We have collected 283 observations for 75 aid recipient countries for which *rent seeking* is available with each country having at most four observations. We will present results by treating all 75 countries as one group (in subsections 4.1. and 4.2) and then partition them into subgroups (in subsection 4.3).

by Svensson’s (2000) point that in practice there is no discrimination between the various forms of rent seeking, as presumably the competing groups equalize marginal costs and benefits between these forms.

4.1. Empirical results

Table 1 presents results from the estimation of the basic model (10a)-(10c). Although this is a system of equations and endogeneity issue is not addressed, OLS estimates displayed in the first column next to each equation are reported to provide results comparable to those of the related literature. The OLS estimates pinpoint a positive sign for *aid* and a negative sign for *rent seeking* in the growth equation. However, the effect of *aid* on *rent seeking* is insignificant in the rent seeking equation. Some other effects should also be noticed. The coefficients of *ldon* and *lpop* in the aid equation have the right signs and are significant at the 5% level. In fact, the hypothesis that the coefficients on *lgdp* and *ldon* are equal and of opposite sign, as predicted by (8a), cannot be rejected by a standard F-test at the 1% significance level. Thus, the redistribution rule assumed in (8a) is not inconsistent with the data. The coefficient on the *government size* is significantly negative in the growth equation, and positive (though insignificant) in the *rent seeking* equation. Concerning control variables, their coefficients have the expected signs and are significant in most cases. For instance, *ethnic tensions* and *assassinations* are significant in the *rent seeking* equation with positive signs. Also, *lgdp* enters with a statistically significant negative sign in all equations. The dummies for Sub-Saharan and East-Asian countries are also significant with expected signs in the growth and rent seeking equations.

Next, the 2SLS estimations are reported in the second column next to each specification. Observe that now, as predicted by the theory, *aid* exerts a positive and significant effect on *rent seeking*.²⁶ Also, the coefficients on *aid* and *rent seeking* in the growth equation have the expected signs and are significant. Combining estimation results,

²⁶ See also Alesina and Weder (2002, p. 1135) who argue that "...an increase in aid is associated with an increase in corruption and vice versa...". Moreover, the lack of statistical significance of *rent seeking* in the aid equation confirms the finding by the same authors and Svensson (2000) that there is no evidence that aid is allocated to countries with less corruption. The combined evidence reinforces the view that the causation runs from aid transfers to rent-seeking activities, which is further strengthened by the insignificance of *aid* in the rent seeking equation when aid enters as an exogenous variable (see the results from the OLS specification).

the coefficient on *aid* implies that a rise of aid transfers as a percentage of GDP by one point would raise the *growth rate* by around 0.5 points in the absence of rent-seeking activities. However, rent-seeking activities also rise as a result of higher aid, so that the net effect of aid on the growth rate is eventually smaller: a rise of aid-to-GDP ratio by one percentage point raises the growth rate by 0.3 percentage points only,²⁷ when the adverse effect of rent seeking is also taken into account.

Regarding the size of the government and its impact on growth and rent-seeking, *government size* enters with a statistically significant negative sign in the *growth* equation with a coefficient value of -0.108, which is not far to that obtained by other growth studies; in other words, *ceteris paribus*, an increase in the size of government by 1 percentage point of GDP will reduce the growth rate on impact by roughly 0.11 percentage points on an annual basis.²⁸ By contrast, the coefficient on *government size* in the rent seeking equation is significantly negative. Although this coefficient turns insignificant in all subsequent specifications (see below), this is somewhat against our theoretical predictions. A possible explanation could be the inclusion of *lgdp* in the estimated equation; Mauro (1998) shows that when one controls for initial GDP, the significance of public spending (with the exception of education expenditures) in explaining corruption is reduced substantially. Finally, observe that, relative to the OLS estimates, the estimated coefficients on control variables and regional dummies retain their signs and significance levels, with the exception of the dummy for East-Asian countries in the *growth* equation, which is now insignificant (see also Burnside and Dollar, 2000).

It is to be noted that the identifying assumptions often raise a number of controversies. We have therefore attempted to tackle with this issue in several ways. As a first step, we test for over-identifying restrictions in all equations; the results are reported in the last row of

²⁷ The net effect is obtained by subtracting the product of the coefficients on *rent seeking* (in the growth equation) and *aid* (in the rent seeking equation) from the coefficient on *aid* in the growth equation.

Table 1 and indicate that the null hypothesis on the validity of the exclusion restrictions cannot be rejected. Second, we conducted several experiments (not reported here) with the variables involving the identifying restrictions. For instance, we included variables like the *Central America*, *Egypt* and the *Franc Zone* dummies, which are assumed to drive aid but not directly growth, in the *growth* equation and we tested their joint significance. In all cases examined the hypothesis that the corresponding coefficients equal zero was soundly accepted.

Finally, the system (10a)-(10c) was estimated by using 3SLS to account for non-zero correlations in the unexplained parts of the regressions (results with the common set of instruments are reported in the third column next to each specification). All coefficients remain roughly unchanged, with the exception of the coefficient on *government size* in the rent seeking equation, which becomes insignificant. The tests for the over-identifying restrictions again suggest that the null on the validity of these restrictions cannot be rejected.²⁹

Regarding the coefficients of main interest, a rise in *aid* again leads to an increase in *rent seeking* (with a somewhat larger coefficient now), whereas *aid* and *rent seeking* appear with similar (though slightly larger) coefficients in the growth equation. The net effect, however, of a rise in the aid-to-GDP ratio by one percentage point on the growth rate remains remarkably close to the one found with 2SLS, indicating that the overall picture is confirmed when the equations are estimated jointly, as long as aid and rent-seeking activities are being treated endogenously.

4.2. Robustness and extensions

We have conducted several robustness tests, all of which confirm the above multivariate relationship between *growth rate*, *rent seeking* and *aid*. Apart from the robustness checks reported in the next paragraphs, the tests also involved the use of alternative definitions for

²⁸ For instance, Barro (2001) reports a corresponding coefficient of -0.157 in his equation for the growth rate.

the variables measuring *aid* and *rent seeking*. For instance, we experimented with the two other measures of aid available by the World Bank, namely Bilateral EDA and Multilateral EDA (see Chang, 1998, for the description of these variables) without any significant changes in the results. We also used the property rights index (available from ICRG), which is the sum of rule of law and expropriation risk, as suggested by Knack and Keefer (1995), and the results remained similar. Finally, we tested the robustness of our results for the possible effect of outliers by dropping one suspected country at a time. None of the resulting regressions for each equation is significantly different from the regressions presented in the paper after the exclusion of the following list of countries (observations), which had the highest residuals: Jordan (1975, 1980), Zaire (all observations), Philippines (all observations), Botswana (1975), Uganda (1980), Gabon: (1975, 1980), Guinea-Bissau (1975).

Table 2 reports a set of these tests (estimated by 3SLS), which refer to some plausible cases. The changes involve the inclusion of *assassinations* and *ethnic* in the growth equation, and the inclusion of *openness* in all three equations. The first modification stems from Burnside and Dollar (2000) and aims at capturing long-term characteristics affecting growth, civil status and policies. The second modification allows for the widely established positive correlation between various measures of openness and economic growth. These variables are often included in empirical growth and aid equations.

In the first column of Table 2, both *assassinations* and *ethnic* are found to be insignificant in the growth equation, in line with the findings by Burnside and Dollar (2000). In the second column, the coefficient on *openness* is significant in the growth equation, but insignificant in the rent seeking equation. Results do not change when all three variables are included in the growth equation.

²⁹ The test statistics are 4.54, 2.46 and 13.62 for the three equations respectively. The corresponding critical value (5%) for the χ^2 distribution is 21.02.

Interestingly, estimates of the coefficient on *aid* in the growth equation become less significant when *openness* is included, whereas the coefficient on *aid* in the rent seeking equation becomes larger in magnitude (and always significant at the 5% level). Consequently, the net effect of *aid* upon *growth* is now close to zero. It thus appears that differences in the degree of openness among aid-recipient countries can be an important determinant of the impact of aid flows on rent-seeking activities. A potential explanation offered by Neeman et al. (2003) is that, in today's open economies, corrupt agents can transfer stolen money outside of the country, thus depleting their country's capital stock.

4.3. Do aid-recipient countries differ?

So far we have treated all 75 aid-recipient countries as a single group. Recall however that the theoretical model distinguishes between those countries that receive relatively large aid transfers and/or have large public sectors and those countries that receive relatively small aid transfers and/or have small public sectors. Only in the former case, *aid* and *government size* can trigger rent-seeking activities. In the latter case, rent-seeking activities (if any) are expected to be independent of *aid* and *government size*, so that only the direct positive effect from aid upon growth should be present.

To test for such differences across countries, we have to partition them on the grounds of the amount of aid they receive, or the size of their public sector, or a combination of both. We have experimented with various definitions of “small” and “large” aid transfers in the data (namely, with different critical values of *aid* below which it would be possible that *rent seeking* is not affected by *aid*) but the data do not distinguish different groups. We believe that this can be justified by the nature of the countries included in the dataset, all of which are “large” aid recipients. Therefore, we do not attempt to partition countries into groups based on the amount of aid received.

The empirical analysis becomes more interesting when we partition countries according to the size of their public sector. Apart from the theoretical prediction in section 2 above, this classification can be motivated by the literature initiated by Burnside and Dollar (2000) and the recent studies of e.g. Collier and Dollar (2004) and Mosley et al. (2004), who point out that aid appears to have a positive impact only in countries with ‘sound’ policies that to a large extent implies small public deficits.³⁰ Table 3 checks the suggested hypothesis of the model for these two groups of countries where, for comparison purposes, we adopt both the core specification of Table 1, as well as the extended specification of Table 2 that includes *assassinations*, *ethnic* and *openness*. These are reported in the first and second column under “large” and “small” government using 121 and 162 observations respectively, with the average government size of the sample taken as the breaking point. In line with the theoretical predictions, the empirical results indicate that, in countries with “large” public sectors, *aid* affects *rent seeking* positively; the coefficients of *aid* are significant in both specifications, although somewhat lower in magnitude compared to the ones reported for the full sample. In contrast, this effect evaporates in countries with “small” public sectors, where the coefficients are insignificant. This implies that the above established adverse effect of foreign aid on incentives is mainly driven by countries with larger-than-average public sectors.³¹

Turning to the growth equation, *rent seeking* affects *growth* negatively irrespectively of government size. In line with the findings in Tables 1 and 2, the estimated coefficients on *rent seeking* are significant, with those in the core specification being smaller in magnitude than those in the extended specification for both groups of countries. The evidence is slightly

³⁰ Along this line we could include the Burnside and Dollar (2000) “good policy” index as an additional variable in the growth and aid equations of the previous section. However, we do not proceed with the empirical investigation because two of the three determinants of the policy index, namely the degree of openness and the budget surplus, are captured by the existing independent variables leaving only the less significant component of the index (inflation) out.

³¹ Another testing procedure is to investigate the statistical significance of the difference in the estimated coefficients. In our case, however, this test would have low power as the variance of the estimate for the countries with “small” public sectors is large and leads the test to accept the null hypothesis too often. We therefore opt for the comparison of the individual estimates for the two groups of countries.

less strong concerning the growth effect of *aid*: the estimated coefficients are significant for countries with “large” public sectors, while in countries with “small” public sectors, where the *aid* effect is found to be larger, they are only marginally significant (in the core specification at 10% level only).³²

To sum up, the data support the theoretical prediction that countries with large public sectors differ from countries with small public sectors. Evidently, in countries with large public sectors, foreign aid triggers rent seeking competition. Although our findings should be interpreted with some caution, they seem to confirm the common belief that rent seeking triggered by aid transfers takes place via state coffers, and this process is facilitated when the size of public sector is relatively large. We believe that the link among aid, fiscal management and incentives merits further research.

5. Conclusions and policy implications

This paper investigated the interrelationship between growth, rent seeking and foreign aid. The primary aim was to explain the recorded poor performance of aid flows in terms of economic growth. To this end, we constructed a general equilibrium model of endogenous growth in which foreign aid can distort individuals’ incentives by pushing them towards socially destructive activities. This indirect adverse effect can partially offset the direct positive impact of aid. This hypothesis was tested for a cross-section of aid-recipient countries where measures of rent seeking are available. The data support the main prediction: aid is less effective in improving growth when its indirect distortion in incentives is taken into account.

³² This may shed some light in the context of Rodrik’s (1998) finding that a rise in openness is associated with an increase in the size of government. An explanation put forward by Rodrik was that in countries, which are more vulnerable to external shocks, the government sector can mitigate risk by taking command of a larger share of the economy’s resources. Along these lines, the evidence presented in Table 3 may bear a complementary explanation, particularly for aid-recipient countries. Namely, aid increases openness and hence provides more scope for government activities, thus rendering the adverse effect on rent-seeking (and growth) stronger.

We close with some brief policy remarks. Although our findings show the adverse effects of aid on private incentives, this should not be taken as a message against the provision of aid. If there is a policy message, this is against poor domestic public institutions, which facilitate the distortion of individual incentives, and so calls for conditional aid, where conditionality focuses on transparency and good functioning of the public sector in recipient countries. By the same argument, greater weight should be placed in the role of independent institutions and NGOs in delivering aid, thus leaving less room for distorted intervention in the allocation of transferred amounts. All this is consistent with the ongoing discussion about the effectiveness of aid under good policies and the attempt by donors to by-pass corrupt governments in order to boost growth in aid-recipient countries.³³

³³ See also e.g. Svensson (2003), Fischer (2003), Easterly et al. (2004), Burnside and Dollar (2004), Mosley et al. (2004) and the wide-ranging discussion in Collier and Dollar (2004).

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TABLE 1. Estimates of the aid impact on growth and rent-seeking: core model (1975-1995, 283 observations)

Dep. variable: <i>growth rate</i>	<i>OLS</i>	<i>2SLS</i>	<i>3SLS</i>	Dep. variable: <i>rent-seeking</i>	<i>OLS</i>	<i>2SLS</i>	<i>3SLS</i>	Dep. variable: <i>aid</i>	<i>OLS</i>	<i>2SLS</i>	<i>3SLS</i>
<i>constant</i>	11.540** (3.61)	11.673** (3.08)	11.740** (3.14)	<i>constant</i>	16.489** (4.14)	10.464** (2.25)	9.890** (2.18)	<i>constant</i>	9.879 (1.10)	9.730 (1.08)	10.541 (1.27)
<i>lgdp</i>	-1.508** (-3.51)	-1.742** (-2.67)	-1.863** (-2.90)	<i>lgdp</i>	-4.676** (-9.48)	-3.767** (-6.30)	-3.863** (-6.64)	<i>lgdp</i>	-1.799** (-9.28)	-1.695** (-6.19)	-1.649** (-6.16)
<i>Sub-Saharan</i>	-3.499** (-5.82)	-4.020** (-5.19)	-4.369** (-5.76)	<i>Sub-Saharan</i>	-3.536** (-4.63)	-4.102** (-4.91)	-4.763** (-5.98)	<i>ipop</i>	-0.818** (-9.08)	-0.808** (-8.77)	-0.797** (-8.96)
<i>East Asia</i>	1.894** (2.58)	1.568* (1.72)	1.452 (1.63)	<i>East Asia</i>	-5.123** (-5.61)	-4.668** (-4.73)	-4.549** (-4.90)	<i>Franc Zone</i>	-0.114 (-0.31)	-0.088 (-0.24)	-0.020 (-0.06)
				<i>Central America</i>	2.378** (2.91)	2.440** (2.79)	2.337** (2.77)	<i>Central America</i>	-0.641* (-1.74)	-0.736* (-1.80)	-0.797** (-2.00)
				<i>assassinations</i>	7.650** (3.99)	7.470** (3.64)	7.084** (3.72)	<i>Egypt</i>	1.036 (1.03)	1.016 (1.00)	0.794 (0.85)
				<i>ethnic</i>	1.009** (4.28)	1.209** (4.65)	1.046** (4.30)	<i>ldon</i>	1.892** (2.03)	1.867** (2.00)	1.748** (2.03)
<i>government size</i>	-0.086** (-3.93)	-0.108** (-3.85)	-0.098** (-3.58)	<i>government size</i>	-0.013 (-0.49)	-0.085** (-2.30)	-0.045 (-1.29)				
<i>aid</i>	0.275** (2.85)	0.493** (2.43)	0.549** (2.75)	<i>aid</i>	-0.003 (-0.03)	0.785** (2.84)	0.876** (3.29)				
<i>rent-seeking</i>	-0.148** (-3.44)	-0.229** (-2.25)	-0.257** (-2.57)					<i>rent-seeking</i>	-0.037 (-1.52)	-0.014 (-0.28)	-0.006 (-0.14)
OR	-	Yes**	Yes**		-	Yes**	Yes**		-	Yes**	Yes**

Notes: *t*-ratios (robust to heteroskedasticity) are in parentheses. An asterisk denotes significance at the 10% level and two asterisks at the 5% level. OR denotes the test of overidentifying restrictions with * denoting acceptance of the null hypothesis at the 5% significance level.

TABLE 2. Estimates of the aid impact on growth and rent-seeking: robustness checks (3SLS, 1975-1995, 283 observations)

Dep. variable: <i>growth rate</i>			Dep. variable: <i>rent-seeking</i>			Dep. variable: <i>aid</i>					
<i>constant</i>	14.150** (2.91)	16.276** (4.10)	18.442** (3.64)	<i>constant</i>	9.849** (2.17)	8.027 (1.64)	8.009 (1.63)	<i>constant</i>	10.756 (1.29)	9.420 (1.13)	9.483 (1.14)
<i>lgdp</i>	-2.800** (-2.42)	-2.678** (-3.86)	-3.608** (-3.06)	<i>lgdp</i>	-3.842** (-6.60)	-3.566** (-5.43)	-3.550** (-5.39)	<i>lgdp</i>	-1.675** (-6.24)	-1.648** (-6.11)	-1.654** (-6.13)
<i>Sub-Saharan</i>	-5.895** (-4.13)	-4.772** (-6.07)	-6.239** (-4.36)	<i>Sub-Saharan</i>	-4.817** (-6.05)	-4.424** (-4.72)	-4.740** (-5.82)	<i>lpop</i>	-0.799** (-8.98)	-0.770** (-7.06)	-0.770** (-7.06)
<i>East Asia</i>	-0.004 (-0.00)	0.827 (0.91)	-0.573 (-0.35)	<i>East Asia</i>	-4.600** (-4.95)	-4.685** (-5.76)	-4.475** (-4.77)	<i>Franc Zone</i>	-0.021 (-0.06)	-0.004 (-0.01)	-0.003 (-0.01)
<i>assassinations</i>	2.152 (0.73)	-	2.647 (0.89)	<i>assassinations</i>	7.153** (3.73)	6.930** (3.68)	7.131** (3.72)	<i>Egypt</i>	0.792 (0.85)	0.684 (0.73)	0.688 (0.74)
<i>ethnic</i>	0.410 (1.13)	-	0.370 (1.01)	<i>ethnic</i>	1.068** (4.37)	1.105** (4.37)	1.120** (4.37)	<i>ldonor</i>	1.735** (2.02)	1.804** (2.11)	1.800** (2.10)
				<i>Central America</i>	2.176** (2.57)	2.481** (2.91)	2.275** (2.65)	<i>Central America</i>	-0.727* (-1.81)	-0.740* (-1.83)	-0.723* (-1.79)
<i>openness</i>		0.020** (2.84)	0.018** (2.20)	<i>openness</i>		-0.009 (-0.97)	-0.009 (-0.96)	<i>openness</i>		0.001 (0.43)	0.001 (0.43)
<i>government size</i>	-0.115** (-3.08)	-0.109** (-3.98)	-0.125** (-3.29)	<i>government size</i>	-0.046 (-1.32)	-0.044 (-1.27)	-0.045 (-1.29)				
<i>aid</i>	0.827** (2.73)	0.283 (1.34)	0.585* (1.80)	<i>aid</i>	0.884** (3.32)	1.028** (3.43)	1.033** (3.45)				
<i>rent-seeking</i>	-0.532* (-1.95)	-0.317** (-3.12)	-0.596** (-2.19)					<i>rent-seeking</i>	-0.011 (-0.23)	-0.006 (-0.13)	-0.007 (-0.16)
OR	Yes**	Yes**	Yes**		Yes**	Yes**	Yes**		Yes**	Yes**	Yes**

Notes: *t*-ratios(robust to heteroskedasticity) are in parentheses. An asterisk denotes significance at the 10% level and two asterisks at the 5% level. OR denotes the test of overidentifying restrictions with * denoting acceptance of the null hypothesis at the 5% significance level.

TABLE 3. Estimates of the aid impact on growth and rent-seeking for ‘large’ (121 obs.) and ‘small’ (162 obs.) government size (3SLS, 1975-1995)

Dep. variable:	“large” government				“small” government				Dep. variable:	“large” government				“small” government			
<i>growth rate</i>									<i>rent-seeking</i>								
<i>constant</i>	11.031**	14.791**	8.975	16.366**	<i>constant</i>	5.602	6.144	12.369**	8.585	<i>constant</i>	31.532	29.398	0.522	0.648			
	(2.20)	(2.14)	(1.15)	(2.17)		(0.88)	(0.88)	(1.26)	(0.92)		(1.64)	(1.50)	(0.12)	(0.15)			
<i>lgdp</i>	-1.797**	-3.606**	-1.745	-3.020**	<i>lgdp</i>	-3.588**	-3.637**	-4.184**	-3.652**	<i>lgdp</i>	-1.796**	-1.855**	-1.360**	-1.453**			
	(-2.31)	(-2.47)	(-1.47)	(-2.39)		(-4.35)	(-3.75)	(-3.28)	(-3.01)		(-3.21)	(-3.26)	(-8.85)	(-8.73)			
<i>Sub-Saharan</i>	-2.899**	-5.537**	-5.237**	-6.668**	<i>Sub-Saharan</i>	-2.957**	-6.171**	-5.876**	-5.315**	<i>ipop</i>	-1.044**	-0.970**	-0.373**	-0.429**			
	(-3.96)	(-3.93)	(-4.18)	(-3.58)		(-2.58)	(-4.38)	(-4.54)	(-3.84)		(-5.93)	(-4.42)	(-6.41)	(-6.69)			
<i>East Asia</i>	-0.016	-3.611	2.047	1.255	<i>East Asia</i>	-5.972**	-6.171**	-3.774**	-3.414**	<i>Franc Zone</i>	0.444	0.544	0.206	0.177			
	(-0.01)	(-1.37)	(1.78)	(0.98)		(-4.21)	(-4.38)	(-3.01)	(-2.72)		(0.51)	(0.61)	(1.16)	(0.99)			
<i>assassinations</i>	-	4.254	-	1.884	<i>assassinations</i>	8.827**	8.820**	5.584*	4.885**	<i>Egypt</i>	-	-	0.848**	0.957**			
		(1.07)		(0.65)		(3.77)	(3.72)	(1.86)	(1.59)				(2.20)	(2.43)			
<i>ethnic</i>	-	0.815	-	0.153	<i>ethnic</i>	1.582**	1.569**	0.749**	0.872**	<i>ldonor</i>	0.075	0.151	1.766**	1.915**			
		(1.45)		(0.43)		(4.50)	(4.00)	(2.12)	(2.46)		(0.04)	(0.08)	(3.75)	(4.02)			
					<i>Central America</i>	3.601**	3.236**	2.211*	2.675**	<i>Central America</i>	-1.425*	-1.227	-0.116	-0.056			
						(2.82)	(2.63)	(1.92)	(2.29)		(-1.79)	(-1.45)	(-0.50)	(-0.24)			
<i>openness</i>	-	0.014	-	0.030**	<i>openness</i>	-	0.005	-	-0.020	<i>openness</i>	-	0.004	-	-0.005**			
		(1.03)		(2.91)			(0.41)		(-1.47)			(0.53)		(-2.09)			
<i>government size</i>	-0.085**	-0.082*	-0.114	-0.197**	<i>government size</i>	0.057	0.0454	-0.083	-0.039								
	(-2.03)	(-1.66)	(-1.43)	(-2.14)		(1.11)	(0.89)	(-0.80)	(-0.35)								
<i>aid</i>	0.324*	0.447**	1.283*	1.038	<i>aid</i>	0.571**	0.490*	1.301	1.513								
	(1.76)	(1.54)	(1.85)	(1.24)		(2.47)	(1.76)	(1.30)	(1.61)								
<i>rent-seeking</i>	-0.266**	-0.762**	-0.321**	-0.460*						<i>rent-seeking</i>	-0.038	-0.050	-0.012	-0.027			
	(-2.25)	(-2.31)	(-2.24)	(-1.93)							(-0.43)	(-0.55)	(-0.44)	(-0.93)			
OR	Yes**	Yes**	Yes**	Yes**		Yes**	Yes**	Yes**	Yes**		Yes**	Yes**	Yes**	Yes**			

Notes: *t*-ratios (robust to heteroskedasticity) are in parentheses. An asterisk denotes significance at the 10% level and two asterisks at the 5% level. OR denotes the test of overidentifying restrictions with * denoting acceptance of the null hypothesis at the 5% significance level.